

## 1. Introduction

(parameter) 가 (assumption) 가  
 가 가 . 가  
 (rank) , 가  
 (median) 가 p-value  
 distribution free, assumption free, statistical inference based on ranks

### 1.1. Nonparametric?

John Arbuthnot (1710) 1942 Wolfowitz가  
 , ,  
 가

#### 1.1.1. advantage ( )

- 가 가 .
- , .
- (rank) .
- 가 .

#### 1.1.2. disadvantage ( )

- 가 .
- 가 .

### 1.2. statistical terms

#### 1.2.1 descriptive and inferential ( )

(descriptive) (statistic)  
 (parameter) (inferential) .

#### 1.2.2. population and sample ( )

(population)  
 (sample) .

, 2000 1 ...

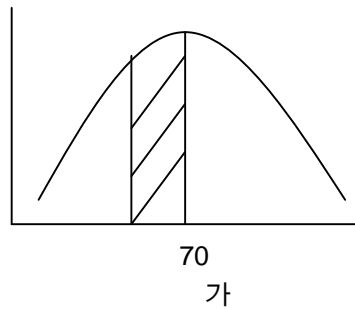
(sample)  $n \leq 20$  가 ( , )

**1.2.3. parameter and statistic ( )**

(parameter)  $\mu$ ,  
 (unknown)  
 $\sigma$ ,  $\rho$   
 (statistic) (estimation)  
 $\bar{x}$ , s 가  
 (median) rank

**1.2.4. random variable and probability density function ( )**

(simple random sample) , (random sample)  
 가 가 ( )  
 가 60 70 가 10



( )

**1.2.5. measurement and categorical ( )**

(measurable numerical)  
 (categorical)  
 , 9 10 A , IQ,  
 가  
 (continuous) , IQ (discrete)

100(kg) 50 (interval) (ratio) 30 15  
 가 .  
 0 가 .  
 가 , 가  
 Likard (5 ) .  
 , , , , , , ,  
 , 가 가 (ordinal)  
 (nominal) . , , ,  
 A, B, C, D, F , ,  
 가

1.2.6. statistical hypothesis ( 가 )

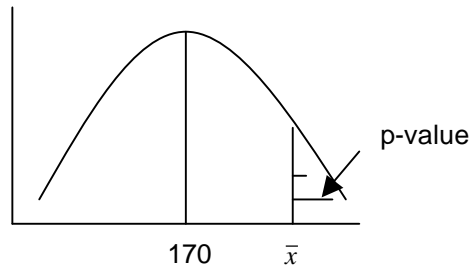
- 가 : 가 (hypothesis)  
 가 .  
 가 가 가 . 가 “ 가 ”,  
 “ ”, “ =100” 가  
 가 가 , 가 .
- : 가 가 ( 가 ) ( 가 )  
 ) 가 .

가 1 (type I error) 2  
 . 가 가 1  
 (significant level)  $\alpha$  . (1- $\alpha$ )  
 (confidence level) .

	가 ( )	가 ( )
가		2 (β)
가	1 (α)	

▪ p-value:

(observed significant level) . p-  
가 가 .



▪ : (1-β)가 (power) . ,  
가 가 .  
가 가 .

$$\{\bar{x} > 120\} \quad \Pr\left(\frac{\bar{x} - u}{s/\sqrt{n}} < \frac{120 - u}{s/\sqrt{n}}\right) \quad (\mu)$$

▪ 가 : 가 .  
가 95% 5% .  
가 가 .  
95% (100, 110) 가  $H_0 : m = 105$  .  
5% . 95% 가 .  
95% 95%

📖 Homework #1 [due 3 8 ]

- 1) 가 170 가 100  
172cm, 10cm .  
5% 가 .  
▪ 가?  
▪ 가 .  
▪ p-value ( 1 ) 가 .  
▪ 172cm 가 가 171, 172,  
173 .

- 171.65cm 가 . 가 171, 172, 173 .
  - 95% .
- 2) 가 20
- 0.6 .
- 가?
  - 가 .
  - p-value 가 .

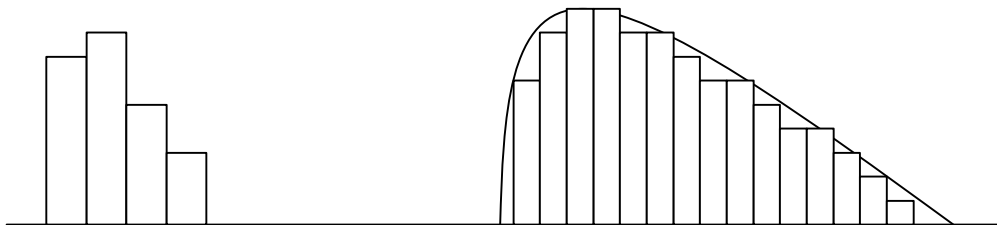
### 1.3. order statistic

### 1.3. order statistic ( )

#### 1.3.1. review

:  
 (random variable) .  $X(c) = x : c =$  ,  $x =$  (real number)  
 ▪ ( ) :  $X(c) = i$  where  $i = 1, 2, 3, 4, 5, 6$

:  
 (probability density function) .  
 ▪ ( 1) :  $f(x) = 1/6$  where  $x = 1, 2, 3, 4, 5, 6$   
 ▪ ( 2)  $f(x) = \frac{1}{\Gamma(r/2)2^{r/2}} x^{(r/2-1)} e^{-x/2}$



: (cumulative distribution function)  $F(x) = \Pr(X \leq x) = \int_{-\infty}^x f(x)dx$

:  $X_1, X_2, \dots, X_n$   
 (random sample) .

#### 1.3.2. definition

가  $n$   $X_1, X_2, \dots, X_n$  가  $X_{(1)}, X_{(2)}, \dots, X_{(n)}$   
 $X_{(1)} < X_{(2)} < \dots < X_{(n)}$  가  $X_{(1)}, X_{(2)}, \dots, X_{(n)}$   
 (order statistic) .

#### 1.3.3.

(minimum)  $X_{(1)}$ , (maximum)  $X_{(n)}$

(range)  $X_{(n)} - X_{(1)}$ , midrange:  $[X_{(1)} + X_{(n)}] / 2$

(median)  $m = [X_{(\frac{n}{2})} + X_{(\frac{n+1}{2})}] / 2$  ( $n$  ),  $m = X_{(\frac{n+1}{2})}$  ( $n$  )

1.3.4.

1)  $X_{(1)}, X_{(2)}, \dots, X_{(n)}$  (joint distribution function)

$$f(x_1, x_2, \dots, x_n) = n! f(x_1) f(x_2) \cdots f(x_n)$$

2)  $X_{(1)}$  marginal distribution function

$$f(x_1) = n[1 - F(x_1)]^{n-1} f(x_1)$$

3)  $X_{(n)}$  marginal distribution function

$$f(x_n) = n[F(x_n)]^{n-1} f(x_n)$$

4)  $n$  (median)  $m$  marginal distribution function

$$f(m) = \frac{(2k)!}{[(k-1)!]^2} \int_m^\infty [F(2m-v)]^{k-1} [1-F(t)]^{k-1} f(2m-t) f(t) dt, n=2k$$



$$f(x) = 2x, 0 < x < 1$$

4

$x_1, x_2, x_3, x_4$

$x_{(1)}, x_{(2)}, x_{(3)}, x_{(4)}$

joint pdf,

$x_{(1)},$

$x_{(n)}$

marginal pdf

1.4.

가? ( )

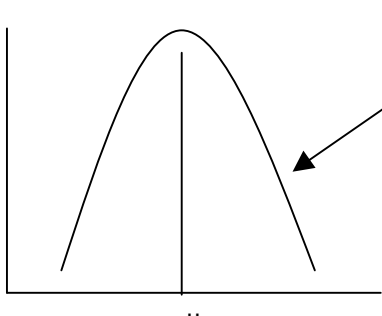
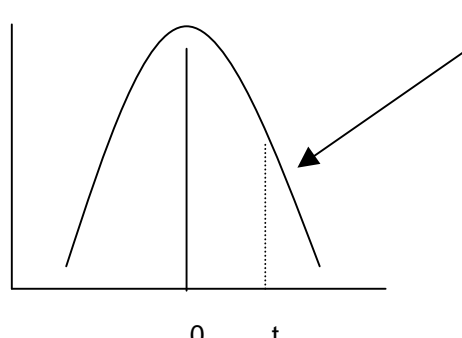
one population + location parameter

60

	60 A 20	가 $f(x)$ 20
	10, 9, 2, 18, 32, 3, 5, 14, 10, 9 6, 8, 1, 15, 12, 9, 4, 7, 8, 3	$x_1, x_2, \dots, x_{20}$
	(elementary statistic) ( ) Location parameter: mean, median	$f(x)$ , 가 parameter
	histogram, stem-leaf, box-plot	$f(x)$ 가

	$f(x)$ skewed, outlier가	
📖	stem-leaf	
	60 가? population mean( $m$ ) median( $M$ )	$\int_{-\infty}^t f(x)dx = 1/2$ t $m, M$
(statistic)	sample mean: $\bar{x} = \sum_{i=1}^{20} x_i / n$ sample median: $[x_{(10)} + x_{(11)}] / 2$	
📖		
	$\hat{m} = \bar{x}, \hat{M} = [x_{(10)} + x_{(11)}] / 2$	
가	60 가 (null hypothesis): $H_0 : m = 9$ 가 (alternative hypothesis): $H_a : m > 9$	9 가? $H_0 : M = 9$ $H_a : M > 9$
가	가 가 ... 가	



<p>( )</p>	<ul style="list-style-type: none"> <li>▪ <math>f(x; u, \sigma)</math></li> <li>▪ [Central Limit Theorem]</li> </ul> <p>normal distribution</p>  <p><math>f_{\bar{x}} \sim Normal(u, \sigma / \sqrt{n})</math></p>
<p>가</p>	<p>가</p>
<p>가 ( )</p>	<p><math>t = \frac{\bar{x} - u}{\sigma / \sqrt{n}} \sim normal(0,1)</math></p> <p>(s)</p> <p>Normal(0,1)</p>  <p>*) 가 가</p>
<p>가 ( )</p>	<ul style="list-style-type: none"> <li>▪ [Student t-distribution]</li> </ul> <p><math>t = \frac{\bar{x} - u}{s / \sqrt{n}} \sim t(n-1)</math></p>
	<p>가 ...</p>

📖 Homework #2 [due 3 15 ]

Simulation

[ ]

1)  $\chi^2$  가 .

2) 20 .

```
data one;
  do i=1 to 20;
    xc2=2*rangam(0, 0.5);
    xc6=2*rangam(0, 5);
    output;
  end;
run;
proc means data=one mean std;
  var xc2 xc6;
run;
```

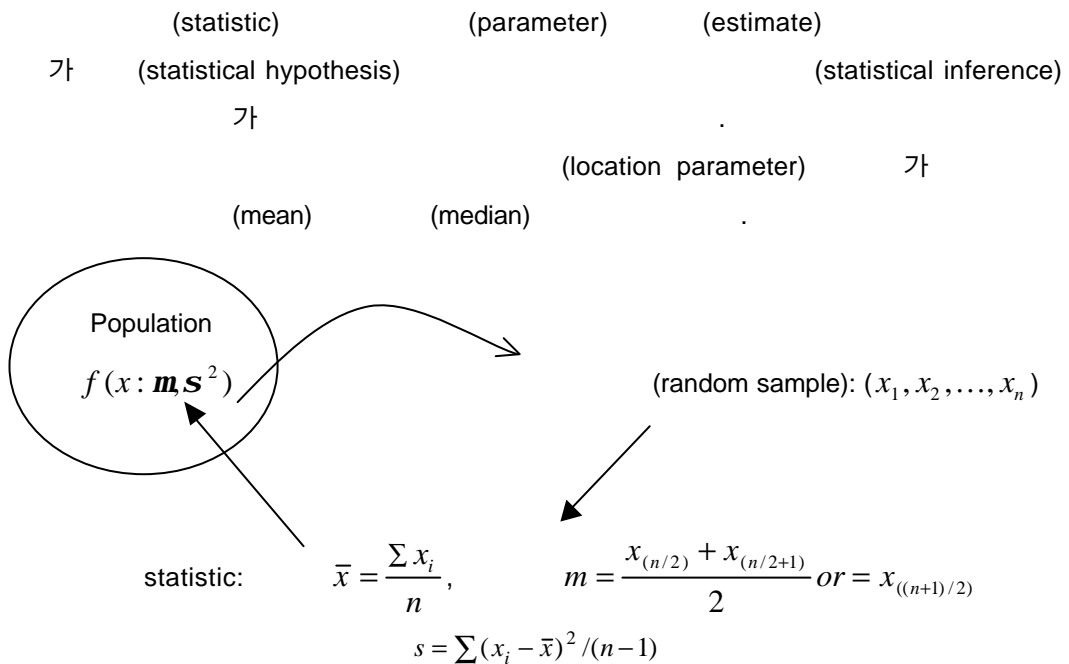
3)  $xc2 \sim \chi^2(\lambda=2)$ ,  $xc6 \sim \chi^2(\lambda=10)$  .

4)  $(xc2, xc6)$  .

5)  $(xc2, xc6)$  2)-4) 1000 .

6)  $(xc2, xc6)$  .

2. One sample

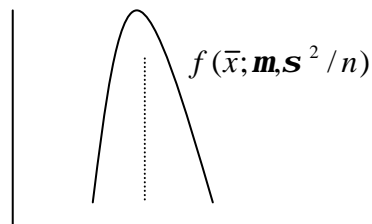
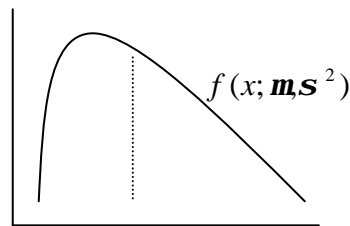


2.1. parametric procedure ( )

mean)

$E(\bar{x}) = \mathbf{m}$ ,  $V(\bar{x}) = \mathbf{s}^2 / n$

(sample



2.1.1. (large sample)

1) (central limit theorem)

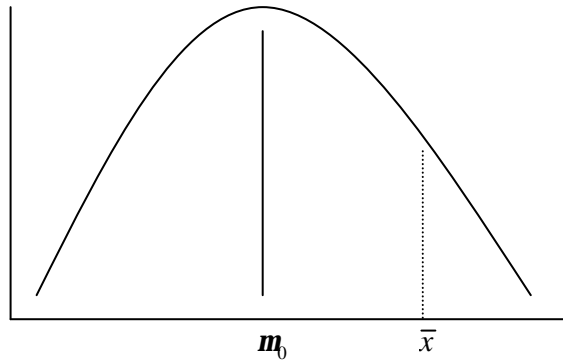
$n$   $\bar{x}$   $f(x)$

2) 가 (hypothesis testing)

- 가 (null hypothesis):  $\mathbf{m} = \mathbf{m}_0$  ( )
- 가 (alternative hypothesis):  $\mathbf{m} \neq \mathbf{m}_0$  (two-sided)  $\mathbf{m} > \mathbf{m}_0$  (one-sided)
- (test statistic): 가 가

( , p-value ) 가

$$T = \frac{\bar{x} - m_0}{s/\sqrt{n}} \sim Normal(0,1)$$



- (conclusion): T (critical region) (p- ) (1 ,  $\alpha$ ) 가 ( 가 ) . two-sided ( ) 가 (  $\alpha$  ) one-sided ( ) 가 가 .

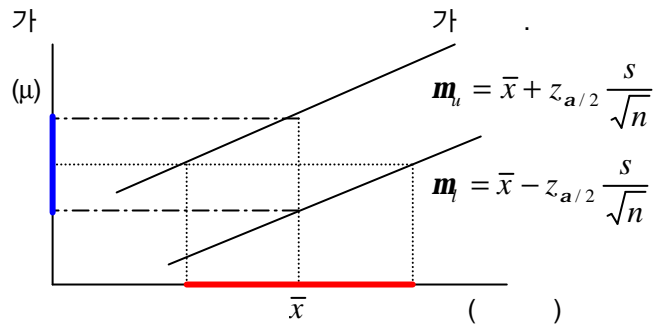
3) (confidence interval)

- 100(1- $\alpha$ )% 
$$\bar{x} \pm z_{\alpha/2} \frac{s}{\sqrt{n}}$$
 (lower limit):  $m_l = \bar{x} - z_{\alpha/2} \frac{s}{\sqrt{n}}$ , (upper limit):  $m_u = \bar{x} + z_{\alpha/2} \frac{s}{\sqrt{n}}$  ( ) 95% 가 가 ( ) 95% 가

... 100

95

4) 100(1- $\alpha$ )%



II Example II      A      (

) 20      (kg)      .

51.0	38.1	44.4	46.4	37.6	46.4	38.3	51.0	38.1	45.1
22.9	40.8	34.9	50.7	68.0	58.0	60.2	38.5	50.7	52.1

A      가 50kg

?

```

class.sas
data one;
  input weight @@;
  cards;
51.0  38.1  44.4  46.4  37.6  46.4  38.3  51.0  38.1  45.1
22.9  40.8  34.9  50.7  68.0  58.0  60.2  38.5  50.7  52.1
;
run;

proc univariate data=one plot;
  var weight;
run;
  
```

Output - (Untitled)

The UNIVARIATE Procedure  
Variable: weight

Moments

N	20	Sum Weights	20
Mean	45.66	Sum Observations	913.2
Std Deviation	10.149379	Variance	103.009895
Skewness	0.09086768	Kurtosis	0.68320875
Uncorrected SS	43653.9	Corrected SS	1957.188
Coeff Variation	22.2281626	Std Error Mean	2.26947014

Basic Statistical Measures

Location		Variability	
Mean	45.66000	Std Deviation	10.14938
Median	45.75000	Variance	103.00989
Mode	38.10000	Range	45.10000
		Interquartile Range	12.80000

NOTE: The mode displayed is the smallest of 4 modes with a count of 2.

Tests for Location: Mu0=0

Test	-Statistic-	-----p Value-----
Student's t	t 20.11923	Pr >  t  <.0001
Sign	M 10	Pr >=  M  <.0001
Signed Rank	S 105	Pr >=  S  <.0001

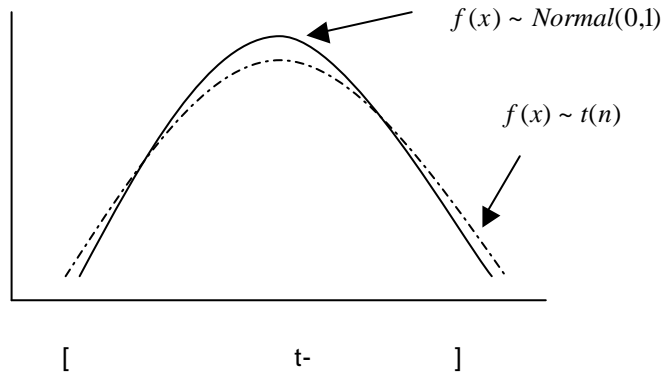
**2.1.2. (small sample)**

1) 가 : n 가  $f(x)$  가

2) t- : 가  $f(x)$  W. S. Gosset

Student

$$\frac{\bar{x} - \mu}{s/\sqrt{n}} \sim t(n-1); \quad =0, \quad =n/(n-2)$$



3)

$$T = \frac{\bar{x} - \mu}{s/\sqrt{n}} \sim t(n-1)$$

$$100(1-\alpha)\% : \bar{x} \pm t(n-1; \alpha/2) \frac{s}{\sqrt{n}}$$

**2.2. Nonparametric procedure I: Sign Test**

n 가

가 (sign)

가

**2.2.1. 가 (hypothesis testing)**

1) 가 (assumption):  $x_1, x_2, \dots, x_n$   $f(x : M)$   
 (random sample) ( )

2) 가 (statistical hypothesis):

- 가 :  $H_0 : M = M_0$
- 가 :  $H_a : M \neq M_0$  ( )  $H_a : M > M_0$   $H_a : M < M_0$  ( )

3) (test statistic):

- $(x_i - M_0)$  .  $M_0$  가
- $x_i = M_0$  가
- 가 +, -
- + 가 ( - 가 ) 가

4) (decision rule):

- $(x_i - M_0)$  “+”, “-“ ( ) p ( ; “+“ ) (Bernoulli)
- “+“ ( “-“ ) (K) (n, p) (Binomial Dist.;  $B(n, p)$ )
- 가  $K \sim B(n, 0.5)$
- “+“ k p-
  - i) k 가  $0.5n$  : p-value =  $\Pr(K \leq k | n, 0.5)$
  - ii) k 가  $0.5n$  : p-value =  $\Pr(K \geq k | n, 0.5)$
- “+“ “-“ k sign test p- p-value =  $\Pr(K \leq k | n, 0.5)$
- p- 가 가

▶ Sign test (population ratio) 가  
 $(H_0 : p = p_0)$  가  
 $\min(np_0, (n-1)p_0) < 9$

📖 [Homework#3 Due 3 20 ] A 가 0.15 가  
 65 가 9 가

|| EXAMPLE || 11 ( )  
 3.5 가 (  $\alpha=5%$  ) (Applied Nonparametric Statistics, W. Daniel)



- 1.8 3.3 5.65 2.25 2.5 3.5 2.75 3.25 3.1 2.7 3.0
- 가 :  $H_0 : M = 3.5, H_a : M \neq 3.5$
  - :  $(x_i - M_0)$  “-“ =9, “+“ =1, 0  
=1 k 1 .
  - $p\text{-} = \Pr(K \leq 1 | n = 10, p = 0.5) = 0.0108$  .
  - $p < 0.025$  ( $\alpha / 2$ ) ( ) 가  
3.5 . k “+“ 3.5

- ▶  $(H_a : M < 3.5)$  p- 0.05 3.5 가?
- ▶ **Large-sample approximation:** 가 12 p-value (normal approximation to the binomial) 가 .

$$z = \frac{(K + 0.5) - 0.5n}{\sqrt{0.5^2 n}} \sim Normal(0,1)$$

p- , z=-2.21  
 $\Pr(K \leq 1 | n = 10, p = 0.5) \approx 0.0136$  . 가 .  
 n=20 n  
 z-

- ▶ **SAS** : UNIVARIATE procedure .

**2.2.2. (confidence intervals)**

- ( : point estimate)  
 가 ( ) .  
 ( : interval estimate) .  
 Sign Test 100(1- $\alpha$ )%  
 .
- $\Pr(K \leq K^*) \leq \alpha/2$  가 .
  - $M_l = X_{(k^*+1)}$  lower limit  $M_u = X_{(n-k^*+1)}$  upper limit .



- $100(1-\alpha)\%$  가 .
- II EXAMPLE II 16 ( ) .
- 95% . (Applied Nonparametric Statistics, W. Daniel)
- 1.90 3.08 9.1 3.53 1.99 3.1 10.16 0.69
- 1.74 2.41 4.01 3.71 8.11 8.23 0.07 3.07

- $\Pr(K \leq 3 | n = 16, p = 0.5) = 0.0105$  ,  $\Pr(K \leq 4 | n = 16, p = 0.5) = 0.0383$
- $(4+1) \alpha = 0.0383 * 2 = 0.0766$  92.34% .
- 92.34%  $(X_{(5)} = 1.99, X_{(12)} = 4.01)$  .

▶ Large-sample approximation : 가 12  
(normal approximation to the binomial)

$$(K^* + 1) \approx (n/2) + z_{\alpha/2} \sqrt{n/4}$$

$$(K^* + 1) \approx (16/2) + 1.96 \sqrt{16/4} \approx 4$$

95%  $(X_{(4)} = 1.9, X_{(13)} = 8.11)$  .

95%  
 $\alpha = 0.0105 * 2 = 0.021$  95%가 97.9% .

📖 [Homework#3 Due 3 20 ] (Applied Nonparametric Statistics, W. Daniel).

**2.1** Lenzer et al. (E2) reported the endurance scores of animals during a 48-hour session of discrimination responding. The median score for animals with electrodes implanted in the hypothalamus was 97.5. Suppose that the experiment was duplicated in another laboratory, except that electrodes were implanted in the forebrain of 12 animals. Assume that investigators observed the endurance scores shown in Table 2.2.

Use the one-sample sign test to see whether the investigators may conclude at the 0.05 level of significance that the median endurance score of animals with electrodes implanted in the forebrain is less than 97.5. What is the *P* value for this test?

**TABLE 2.2**

**Endurance scores of animals with electrodes implanted in forebrain**

---

93.6	89.1	97.7	84.4	97.8	94.5	88.3	97.5	83.7	94.6	85.5	82.6
------	------	------	------	------	------	------	------	------	------	------	------

- 2.2 Iwamoto (E3) found that the mean weight of a sample of a particular species of adult female monkey from a certain locality was 8.41 kg. Suppose that a sample of adult females of the same species from another locality yielded the weights shown in Table 2.3.

Can we conclude that the median weight of the population from which this second sample was drawn is greater than 8.41 kg? Use the one-sample sign test and a 0.05 level of significance. What is the  $P$  value for this test?

**Weights of female monkeys, kilograms**      **TABLE 2.3**

8.30	9.50	9.60	8.75	8.40	9.10	9.25	9.80	10.05	8.15	10.00	9.60	9.80	9.20	9.30
------	------	------	------	------	------	------	------	-------	------	-------	------	------	------	------

- 2.6 Armstrong (E9) studied the daily exposure, in minutes, of 10 North Hawaiian families to the risk of a motor vehicle accident. Suppose that a similar survey in another area yielded the exposure times shown in Table 2.10. Find the point estimate and construct an approximate 95% confidence interval for the population median.

**TABLE 2.10**

**Exposure times per day, in minutes, of individuals to motor-vehicle accidents**

18.3	45.2	19.1	57.0	63.9	10.3	12.1	35.5	36.6	74.6
10.5	27.8	44.9	40.9	63.7	40.8	59.1	31.5	40.1	8.1

- 2.7 Abu-Ayyash (E10) found that the median education of heads of households living in mobile homes in a certain area was 11.6 years. Suppose that a similar survey conducted in another area revealed the educational levels of heads of households shown in Table 2.11. Find the point estimate, and construct the approximate 95% confidence interval for the population median.

**TABLE 2.11**

**Educational levels (years of school completed) of heads of households residing in mobile homes**

13	6	6	12	12	10	9	11	14	8	7	16	15	8	7
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### 2.3. Nonparametric procedure II: Wilcoxon Signed-ranks test

sign test 가  $(x_i - M_0)$

#### 2.3.1. 가 (hypothesis testing)

1) 가 (assumption):  $x_1, x_2, \dots, x_n$   $f(x : M)$   
 (random sample) ( ) sign test  
 가

2) 가 (statistical hypothesis):

- 가 :  $H_0 : M = M_0$  ( ),  $H_0 : M \leq M_0$   $H_0 : M \geq M_0$  ( )
- 가 :  $H_a : M \neq M_0$  ( ),  $H_a : M > M_0$   $H_a : M < M_0$  ( )

3) (test statistic):

- $D_i = (x_i - M_0)$   $D_i = 0$   
 $M_0$  가
- $D_i$   $|D_i|$   $|D_i|$  가  
 $3$   
 $|D_i|$   $1, 2, 3$   
 $|D_i|$   $(1=2+3)/3=2, 3$

2

- $D_i$
- 가 -  $T_-$ , 가 +  $T_+$   
 $T_+ = [n(n+1)/2] - T_-$  가
- 가  $T_-$   $T_+$   $T_-$   
 $T_+$  가 가
- 가  $H_a : M \neq M_0$   $T = \min(T_+, T_-)$
- 가  $H_a : M > M_0$   $T = T_-$
- 가  $H_a : M < M_0$   $T = T_+$

4) (decision rule):

- Wilcoxon 가 n  
 (critical region)  $\alpha$   
 가 가

▶ Sign test (population ratio) 가  
 $(H_0 : p = p_0)$  가 가

$$\min(np_0, (n-1)p_0) < 9$$

II EXAMPLE II

IQ

107

IQ

15

(  $\alpha=5%$  ) (Applied Nonparametric Statistics, W. Daniel)

99 100 90 94 135 108 107 111 119 104 127 109 117 105 125

가 :  $H_0 : M = 107$  ,  $H_a : M \neq 107$

가 :  $T_+ = 64.5$  ,  $T_- = 40.5$

$$T = \min(64.5, 40.5) = 40.5$$

IQ	$D_i = X_i - M_0$	Rank of $ D_i $	Signed rank of $ D_i $
99	-8	7	-7
100	-7	6	-6
90	-17	11	-11
94	-13	10	-10
135	+28	14	+14
108	+1	1	+1
107	0		Eliminate from analysis
111	+4	5	+5
119	+12	9	+9
104	-3	4	-4
127	+20	13	+13
109	+2	2.5	+2.5
117	+10	8	+8
105	-2	2.5	-2.5
125	+18	12	+12
			$T_+ = 64.5$
			$T_- = 40.5$

Wilcoxon signed-ranks Table

{  $T \leq 21$  }

$T = 40.5$

가

IQ

가  $\alpha = 0.0247 * 2 = 0.494$

$H_0 : M \geq 107$  vs.  $H_a : M < 107$

{  $T \leq 25$  }

$T = T_- = 40.5$  가

가

IQ가

Large-sample approximation:

가 20

Wilcoxon

가

$$T^* = \frac{T - n(n+1)/4}{\sqrt{n(n+1)(2n+1)/24}} \sim Normal(0,1)$$

$T^*$

$T$  ,

$T_-$   $T_+$

z-



▶ sampling distribution of  $T_+$  : 가 n=4  $T_+$

Wilcoxon

+	$T_+$	+	$T_+$	+	$T_+$	+	$T_+$
	0	4	4	2,3	5	1,2,4	7
1	1	1,2	3	2,4	6	1,3,4	8
2	2	1,3	4	3,4	7	2,3,4	9
3	3	1,4	5	1,2,3	6	1,2,3,4	10

Wilcoxon 가 (n=4)

2.3.2. (confidence intervals)

Wilcoxon

1)  $u_{ij} = \frac{x_i + x_j}{2}, 1 \leq i \leq j \leq n$  ?  ${}_{10}C_2 + 10 = 55$

2)  $u_{ij}$

가

3) Wilcoxon n P 가 T .  $K(=T+1)$   
 $u_{ij}$  (lower limit) K dl (upper limit)

|| EXAMPLE || 10 95%

(Applied Nonparametric Statistics, W. Daniel)

28.5 25.2 28.7 41 29.1 21.3 37.7 39.9 26.8 28.8

1)  $u_{ij} = \frac{x_i + x_j}{2}, 1 \leq i \leq j \leq n$

2)  $u_{ij}$

$u_{ij(28)} = 31.45$

3) n=10 P=0.0244 T=8 . (  $u_{ij(9)} = 27.75$  ,  $u_{ij(47)} = 35.05$  )

95.12%

▶ Large-sample approximation : 가 20

Wilcoxon

$$K \approx \frac{n(n+1)}{4} - z_{\alpha/2} \sqrt{\frac{n(n+1)(2n+1)}{24}}$$

 [Homework#4 Due 3 27 ] (Applied Nonparametric Statistics, W. Daniel).

- 2.3** Malina (E5) reported the results of a study of weights of football players at the University of Texas at Austin between 1899 and 1970. Suppose that the weights of a random sample of 15 football players during the past 10 years at another large state university are those shown in Table 2.6.

Can we conclude that the median weight of the population from which this sample was drawn is greater than 163.5 pounds? Let  $\alpha = 0.05$ . What is the  $P$  value for this test?

**Weights of football players** **TABLE 2.6**

Player	Weight	Player	Weight
1	188.0	9	214.4
2	211.2	10	221.0
3	170.8	11	162.0
4	212.4	12	222.8
5	156.9	13	174.1
6	223.1	14	210.3
7	235.9	15	195.2
8	183.9		

- 2.4** Moore and Ogletree (E6) investigated the readiness of pupils at the beginning of the first grade. They compared scores on a readiness test of pupils who had attended a Head Start program for a full year with the scores of those who had not. Suppose that a random sample of 20 pupils who had not attended a Head Start program achieved the scores on the readiness test shown in Table 2.7.

Can we conclude that the median score of the population represented by this sample is less than 45.32? What is the  $P$  value for this test?

**TABLE 2.7**

**Readiness test scores of pupils who did not attend a Head Start program**

Pupil	Readiness score	Pupil	Readiness score
1	33	11	41
2	19	12	31
3	40	13	46
4	35	14	51
5	51	15	34
6	41	16	37
7	27	17	36
8	23	18	55
9	39	19	52
10	21	20	32

[homework #4 ] Sign Test 2.7 Wilcoxon  
 approximation 95% confidence interval . approximation  
 95% , large-sample

**2.4. One-sample runs test for randomness**

가 가 (random sample) 가 . 가  
 가 randomness가 가 Run

**2.4.0. Example**

1) p control chart:

control limit  
 randomness가 pattern control 가

2)

**2.4.1.**

0) :

randomness runs ( 가 ) run .  
 MFMFMFMFMF → runs 10 . pattern  
 MMMMMFFFFFF → runs 2 가 5 .

1) 가

가 n, 가 n<sub>1</sub>,

가 n<sub>2</sub> n=n<sub>1</sub>+n<sub>2</sub> .

2) 가

가 (null hypothesis): randomness .

가 (alternative hypothesis): randomness .

3) : runs (r) . 가 .

4) (decision rule):

- r (n<sub>1</sub>, n<sub>2</sub>) 2 (critical values of r in the runs test)  
 \* ) ( ) 가

II EXAMPLE II

A

+, - random 가

TABLE 2.18

Departures from normal of daily temperatures recorded at Atlanta, Georgia, during November, 1974

<b>Day</b>	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15
<b>Departure from normal</b>	12	13	12	11	5	2	-1	2	-1	3	2	-6	-7	-7	-12
<b>Day</b>	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30
<b>Departure from normal</b>	-9	6	7	10	6	1	1	3	7	-2	-6	-6	-5	-2	-1

Source: Local Climatological Data, U.S. Department of Commerce, National Oceanic and Atmospheric Administration, Environmental Data Service, National Climatic Center, Federal Building, Asheville, N.C., November 1974.

- 가 : 가 random
- 가 : 가 random
- :  $n_1 = 17$  (+),  $n_2 = 13$  (-)  $n = 30$   $r = 8$
- 2 { $r \leq 10$ } { $r \geq 22$ }  $r = 8$   
 가 5%

▶ Large-sample approximation:  $n_1, n_2$  가 20

가

$$T^* = \frac{r - \{[(2n_1n_2)/(n_1 + n_2)] + 1\}}{\sqrt{\frac{2n_1n_2(2n_1n_2 - n_1 - n_2)}{(n_1 + n_2)^2(n_1 + n_2 - 1)}}} \sim Normal(0,1)$$



📖 [Homework#5 Due 3 29 ] (Applied Nonparametric Statistics, W. Daniel).

- 2.19** Table 2.19 shows the actual daily occurrence of sunshine in Atlanta during November 1974, as a percentage of the possible time the sun could have shone if it had not been for cloudy skies. The data are from the U.S. Department of Commerce (E28). Dichotomize the observations according to whether the amount of sunshine was more than 50% of possible or 50% or less, and test the null hypothesis that the pattern of occurrence of the two types of day is random.

**TABLE 2.19** Percentage of day during which sunshine occurred in Atlanta, November, 1974

Day	Percentage	Day	Percentage	Day	Percentage
1	85	11	31	21	87
2	85	12	86	22	100
3	99	13	100	23	100
4	70	14	0	24	88
5	17	15	100	25	50
6	74	16	100	26	100
7	100	17	46	27	100
8	28	18	7	28	100
9	100	19	12	29	48
10	100	20	54	30	0

Source: *Local Climatological Data*, U.S. Department of Commerce, National Oceanic and Atmospheric Administration, Environmental Data Service, National Climatic Center, Federal Building, Asheville, N.C., November 1974.

- 2.22** Littler et al. (E31) studied the blood flow in lung capillaries in 16 patients with scoliosis or neuromuscular weakness. They reported the sex of the patients in the following order:

F F F M F F M M M F F F F F M

Test the null hypothesis that this sequence is random.

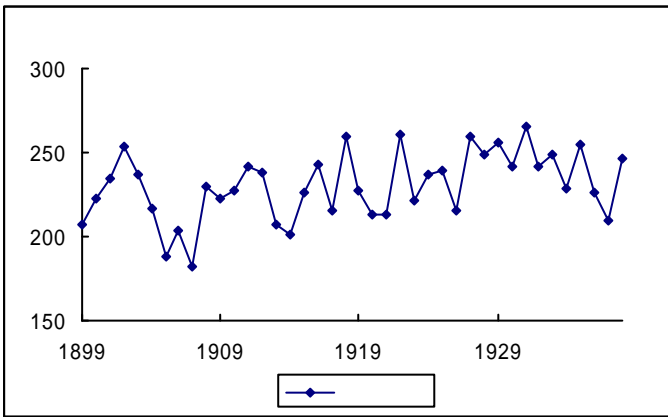
2.5. Cox-Stuart Test for Trend

trend가

Sign test

Cox, D. R. and A. Stuart, "Some Quick Tests for trend in Location and dispersion", Biometrika, 42 (1955), 80-95

2.5.0. Example ( )



년도	수확기간	년도	수확기간
1899	207	1919	227
1900	223	1920	213
1901	235	1921	213
1902	254	1922	261
1903	237	1923	222
1904	217	1924	237
1905	188	1925	239
1906	204	1926	216
1907	182	1927	260
1908	230	1928	249
1909	223	1929	256
1910	227	1930	242
1911	242	1931	266
1912	238	1932	242
1913	207	1933	249
1914	201	1934	228
1915	226	1935	255
1916	243	1936	226
1917	215	1937	209
1918	259	1938	247

2.5.1.

1) 가

2) 가

가 (null hypothesis): trend가

가 (alternative hypothesis)1: upward trend가

가 (alternative hypothesis)2: downward trend가

가 (alternative hypothesis)3: trend가

3) :  $(x_i, x_{c+i})$   $(x_{c+i} - x_i)$  (+, -) trend  
 . n  $c = n/2$  n  $c = (n+1)/2$

가

4) (decision rule):

가 0 + - 가

sign test

II EXAMPLE II

A

- 가 : trend가
- 가 : trend가 . ( )
- : (207, 227), (223,213), ... (n'=40) → + =6 - =14,  
6 . (sign test )
- p-value  $P(K \leq 6 | n = 20, p = 0.5) = 0.0577$  0.025  
가 trend가

▶ Large-sample approximation: sign test

📖 [Homework#5 Due 3 29 ] (Applied Nonparametric Statistics, W. Daniel).

**2.24** The 1972 edition of the *FAA Statistical Handbook of Aviation* (E34) gives the information on annual United States exports of aircraft, aircraft parts, and accessories shown in Table 2.24. Do these data reflect an upward trend in exports? Let  $\alpha = 0.05$ . What is the  $P$  value?

**TABLE 2.24** Number of U.S. aircraft exports, 1947–1971

Year	Aircraft exports*	Year	Aircraft exports*
1947	3125	1960	2336
1948	2258	1961	2459
1949	881	1962	2131
1950	756	1963	2251
1951	894	1964	2577
1952	1180	1965	3129
1953	1377	1966	3611
1954	1053	1967	3881
1955	1714	1968	3682
1956	1711	1969	3322
1957	2025	1970	3383
1958	1689	1971	2904
1959	1628		

Source: *FAA Statistical Handbook of Aviation*, Department of Transportation, Federal Aviation Administration, for sale by the Superintendent of Documents, U.S. Government Printing Office, Washington, D.C., 1972.

\* 1949–1954, civil only